

A Note on Ratio Estimators in two Stage Sampling

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Abstract- In this paper a number of ratio estimators in two stage sampling are considered and their efficiencies are compared with an estimator without use auxiliary information. Numerical illustration is provided to compare the efficiencies.

survey operations even if estimates derived from multi stage sampling are likely to be less efficient than those of the unrestricted uni-stage sampling. Sometimes, all auxiliary variable x compared with y may be available either at primary stage or of secondary stage or at both the stages to in case the efficiency of estimate of the finite population parametric functions such as population means or total. Thus, in the following we consider different ratio estimators of population mean of the study variable y in two –stage sampling using knowledge on a single auxiliary variable x .

I. INTRODUCTION

In large scale sample surveys it is usual practice to adopt multi stage sampling to estimate the population mean or total of the study variable y . The main purpose of using multi stage sampling in place of unrestricted unistage sampling is to reduce the cost of

II. NOTATIONS

Now, consider a finite population U partitioned into N first stage units (fsu) denoted by U_1, U_2, \dots, U_N . Let M_i be the number of second stage units in $U_i (i = 1, 2, \dots, N)$. Define $M = \sum_{i=1}^N M_i$ and $\bar{M} = \frac{1}{N} \sum_{i=1}^N M_i$. Let y_{ij} and x_{ij} denote values of the study variable y and the auxiliary variable x respectively for the j th ssu of $U_i = (j = 1, 2, \dots, M_i; i = 1, 2, \dots, N)$.

Define,
$$\bar{Y}_i = \frac{1}{M_i} \sum_{j=1}^{M_i} y_{ij} \quad \text{and} \quad \bar{X}_i = \frac{1}{M_i} \sum_{j=1}^{M_i} x_{ij} \quad (i = 1, 2, \dots, N)$$

The population mean of y ,
$$\bar{Y} = \frac{1}{N} \sum_{i=1}^N u_i \bar{Y}_i$$
 and the population mean of x ,
$$\bar{X} = \frac{1}{N} \sum_{i=1}^N u_i \bar{X}_i$$
, where $u_i = \frac{M_i}{M}$.

Further, define
$$R = \frac{\bar{Y}}{\bar{X}} \quad \text{and} \quad R_i = \frac{\bar{Y}_i}{\bar{X}_i} \quad (i = 1, 2, \dots, N)$$

$$S_{by}'^2 = \frac{1}{N-1} \sum_{i=1}^N (u_i \bar{Y}_i - \bar{Y})^2$$

$$S_{bx}'^2 = \frac{1}{N-1} \sum_{i=1}^N (u_i \bar{X}_i - \bar{X})^2$$

$$S_{bxy}' = \frac{1}{N-1} \sum_{i=1}^N (u_i \bar{Y}_i - \bar{Y})(u_i \bar{X}_i - \bar{X})$$

$$S_{iy}^2 = \frac{1}{M_i - 1} \sum_{j=1}^{M_i} (y_{ij} - \bar{Y}_i)^2, \quad i = 1, 2, \dots, N.$$

$$S_{ix}^2 = \frac{1}{M_i - 1} \sum_{j=1}^{M_i} (x_{ij} - \bar{X}_i)^2, \quad i = 1, 2, \dots, N.$$

$$S_{ixy} = \frac{1}{M_i - 1} \sum_{j=1}^{M_i} (x_{ij} - \bar{X}_i)(y_{ij} - \bar{Y}_i), \quad i = 1, 2, \dots, N.$$

Define,

$$\bar{y}_i = \frac{1}{m_i} \sum_{j \in S_i} y_{ij} = \frac{1}{m_i} \sum_{j=1}^{m_i} y_{ij}$$

$$\bar{x}_i = \frac{1}{m_i} \sum_{j \in S_i} x_{ij} = \frac{1}{m_i} \sum_{j=1}^{m_i} x_{ij}$$

$$\bar{y} = \frac{1}{n} \sum_{i \in S} u_i \bar{y}_i = \frac{1}{n} \sum_{i=1}^n u_i \bar{y}_i$$

$$\bar{x} = \frac{1}{n} \sum_{i \in S} u_i \bar{x}_i = \frac{1}{n} \sum_{i=1}^n u_i \bar{x}_i$$

$$S_{by}^{\prime 2} = \frac{1}{n-1} \sum_{i=1}^n (u_i \bar{y}_i - \bar{y})^2$$

$$S_{bx}^{\prime 2} = \frac{1}{n-1} \sum_{i=1}^n (u_i \bar{x}_i - \bar{x})^2$$

$$S'_{bxy} = \frac{1}{n-1} \sum_{i=1}^n (u_i \bar{y}_i - \bar{y})(u_i \bar{x}_i - \bar{x})$$

$$S_{iy}^2 = \frac{1}{m_i - 1} \sum_{j=1}^{m_i} (y_{ij} - \bar{y}_i)^2$$

$$S_{ix}^2 = \frac{1}{m_i - 1} \sum_{j=1}^{m_i} (x_{ij} - \bar{x}_i)^2$$

$$S_{ixy} = \frac{1}{m_i - 1} \sum_{j=1}^{m_i} (y_{ij} - \bar{y}_i)(x_{ij} - \bar{x}_i)$$

$$\rho_b = \frac{S'_{bxy}}{S'_{by} S'_{bx}}, \quad \rho_i = \frac{S_{ixy}}{S_{iy} S_{ix}} \quad (i = 1, 2, \dots, N)$$

$$C_{bx} = \frac{S'_{bx}}{\bar{X}}, \quad C_{by} = \frac{S'_{by}}{\bar{Y}}$$

$$C_{ix} = \frac{S_{ix}}{\bar{X}_i}, \quad C_{iy} = \frac{S_{iy}}{\bar{Y}_i}, \quad (i = 1, 2, \dots, N)$$

III. RATIO ESTIMATORS

Consider the following ratio type estimators under two-stage sampling scheme.

$$T_0 = \frac{1}{n} \sum_{i=1}^n u_i \bar{y}_i, \quad \text{without using auxiliary information on } x.$$

$$T_1 = \frac{\frac{1}{n} \sum_{i=1}^n u_i \bar{y}_i}{\frac{1}{n} \sum_{i=1}^n u_i \bar{x}_i} \bar{X}.$$

$$T_2 = \frac{1}{n} \sum_{i=1}^n u_i \frac{\bar{y}_i}{\bar{X}_i} \bar{X}_i$$

$$T_3 = \frac{\frac{1}{n} \sum_{i=1}^n u_i \bar{y}_i}{\frac{1}{n} \sum_{i=1}^n u_i \bar{X}_i} \bar{X}$$

Smith (1969)

$$T_4 = \frac{\frac{1}{n} \sum_{i=1}^n u_i \frac{\bar{y}_i}{\bar{X}_i} \bar{X}_i}{\frac{1}{n} \sum_{i=1}^n u_i \bar{X}_i} \bar{X}$$

Murthy (1967)

$$T_5 = \frac{\frac{1}{n} \sum_{i=1}^n u_i \bar{y}_i}{\frac{1}{n} \sum_{i=1}^n u_i \bar{X}_i} \frac{1}{n} \sum_{i=1}^n u_i \bar{X}_i$$

These estimators belong to the class of estimators considered by Panda (1998).

T_1 requires advance knowledge on \bar{X}_j , T_2 and T_5 require advance knowledge on $\bar{X}_i (i = 1, 2, \dots, n)$; T_3 and T_4 require advance knowledge on \bar{X} and $\bar{X}_i (i = 1, 2, \dots, n)$.

IV. BIASES AND MEAN SQUARE ERRORS OF ESTIMATORS

As known, T_0 is an unbiased estimator of \bar{Y} . $T_1, T_2, T_3, T_4,$ and T_5 are biased estimates of \bar{Y} and upto $O\left(\frac{1}{n}\right)$, the biases are :

$$B(T_1) = \bar{Y} \left[\left(\frac{1}{n} - \frac{1}{N} \right) \left\{ \frac{S'^2_{bx}}{\bar{X}^2} - \frac{S'_{bxy}}{\bar{Y} \bar{X}} \right\} + \frac{1}{nN} \sum u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i} \right) \left\{ \frac{S^2_{ix}}{\bar{X}^2} - \frac{S_{ixy}}{\bar{Y} \bar{X}} \right\} \right]$$

$$B(T_2) = \frac{1}{N} \sum_{i=1}^N u_i \bar{Y}_i \left(\frac{1}{m_i} - \frac{1}{M_i} \right) \left(\frac{S^2_{ix}}{\bar{X}_i^2} - \frac{S_{ixy}}{\bar{X}_i \bar{Y}_i} \right)$$

$$B(T_3) = \bar{Y} \left(\frac{1}{n} - \frac{1}{N} \right) \left[\frac{S'^2_{bx}}{\bar{X}^2} - \frac{S'_{bxy}}{\bar{Y} \bar{X}} \right]$$

$$B(T_4) = \left(\frac{1}{n} - \frac{1}{N} \right) \bar{Y} \left\{ \frac{S'^2_{bx}}{\bar{X}^2} - \frac{S'_{bxy}}{\bar{X}_i \bar{Y}} \right\} - \text{COV} \left[\frac{1}{n} \sum_{i=1}^n u_i \bar{Y}_i \left(\frac{1}{m_i} - \frac{1}{M_i} \right) \left(\frac{S^2_{ix}}{\bar{X}_i^2} - \frac{S_{ixy}}{\bar{X}_i \bar{Y}_i} \right), \frac{1}{n} \sum_{i=1}^n u_i \bar{X}_i \right]$$

$$B(T_5) = \bar{Y} \left[\left(\frac{1}{n} - \frac{1}{N} \right) \frac{S'^2_{bx}}{\bar{X}^2} + \frac{1}{nN} \sum_{i=1}^N u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i} \right) \left\{ \frac{S^2_{ix}}{\bar{X}^2} - \frac{S_{ixy}}{\bar{Y} \bar{X}} \right\} \right]$$

To $O\left(\frac{1}{n}\right)$ the mean square errors (MSE) of T_1, T_2, T_3, T_4 and T_5 are given by

$$\text{MSE}(T_0) = \left(\frac{1}{n} - \frac{1}{N} \right) S'^2_{by} + \frac{1}{nN} \sum_{i=1}^N u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i} \right) S^2_{iy}$$

$$MSE(T_1) = \left(\frac{1}{n} - \frac{1}{N}\right) [S'_{by}{}^2 + R^2 S'_{bx}{}^2 - 2RS'_{xy}] + \frac{1}{nN} \sum_{i=1}^N u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i}\right) (S_{iy}^2 + R^2 S_{ix}^2 - 2RS_{ixy})$$

$$MSE(T_2) = \left(\frac{1}{n} - \frac{1}{N}\right) S'_{by}{}^2 + \frac{1}{nN} \sum_{i=1}^N u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i}\right) (S_{iy}^2 + R_i^2 S_{ix}^2 - 2R_i S_{ixy})$$

$$MSE(T_3) = \left(\frac{1}{n} - \frac{1}{N}\right) [S'_{by}{}^2 + R^2 S'_{bx}{}^2 - 2RS'_{xy}] + \frac{1}{nN} \sum_{i=1}^N u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i}\right) S_{iy}^2$$

$$MSE(T_4) = \left(\frac{1}{n} - \frac{1}{N}\right) [S'_{by}{}^2 + R^2 S'_{bx}{}^2 - 2RS'_{yx}] + \frac{1}{nN} \sum_{i=1}^N u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i}\right) (S_{iy}^2 + R_i^2 S_{ix}^2 - 2R_i S_{ixy})$$

$$MSE(T_5) = \left(\frac{1}{n} - \frac{1}{N}\right) S'_{by}{}^2 + \frac{1}{nN} \sum_{i=1}^N u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i}\right) (S_{iy}^2 + R^2 S_{ix}^2 - 2RS_{ixy})$$

V. COMPARISON OF EFFICIENCIES

The sufficient conditions under which T_1, T_2, T_3, T_4, T_5 would be more efficient than T_0 (without using auxiliary variable) are given in Table 1.

Table 1

| | |
|--|---|
| $T_1 : \rho_b > \frac{1}{2} \frac{C_{bx}}{C_{by}}$ | $\rho_i > \frac{1}{2} \left(\frac{R}{R_i}\right) \frac{C_{ix}}{C_{iy}}$ for all $i (i= 1, 2, \dots, N)$ |
| $T_2 : \text{-----}$ | $\rho_i > \frac{1}{2} \frac{C_{ix}}{C_{iy}}$ |
| $T_3 : \rho_b > \frac{1}{2} \frac{C_{bx}}{C_{by}}$ | ----- |
| $T_4 : \rho_b > \frac{1}{2} \frac{C_{bx}}{C_{by}}$ | $\rho_i > \frac{1}{2} \frac{C_{ix}}{C_{iy}}$ for all $i. (i = 1, 2, \dots, N)$ |
| $T_5 : \text{-----}$ | $\rho_i > \frac{1}{2} \left(\frac{R}{R_i}\right) \frac{C_{ix}}{C_{iy}}$ for all $i. (i = 1, 2, \dots, N)$ |

VI. ESTIMATION OF VARIANCES

We have

$$Est(S'_{bxy}) = s'_{bxy} - \frac{1}{n} \sum_{i=1}^n u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i}\right) s_{ixy}$$

$$Est(S'_{bx}{}^2) = s'_{bx}{}^2 - \frac{1}{n} \sum_{i=1}^n u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i}\right) s_{ix}^2$$

$$Est(S'_{by}{}^2) = s'_{by}{}^2 - \frac{1}{n} \sum_{i=1}^n u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i}\right) s_{iy}^2$$

$$\text{Est} \left[\frac{1}{N} \sum_{i=1}^N u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i} \right) S_{ixy} \right] = \frac{1}{n} \sum_{i=1}^n u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i} \right) s_{ixy}$$

$$\text{Est} \left[\frac{1}{N} \sum_{i=1}^N u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i} \right) S_{ix}^2 \right] = \frac{1}{n} \sum_{i=1}^n u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i} \right) s_{ix}^2$$

$$\text{Est} \left[\frac{1}{N} \sum_{i=1}^N u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i} \right) S_{iy}^2 \right] = \frac{1}{n} \sum_{i=1}^n u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i} \right) s_{iy}^2$$

Thus, Est MSE(T₁)

$$\begin{aligned} &= \left(\frac{1}{n} - \frac{1}{N} \right) \{ s_{by}'^2 + \hat{R}^2 s_{bx}'^2 - 2\hat{R}s'_{bxy} \} \\ &\quad + \frac{1}{nN} \sum_{i=1}^n u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i} \right) \{ s_{iy}^2 + \hat{R}^2 s_{ix}^2 - 2\hat{R}s_{ixy} \} \end{aligned}$$

Est MSE (T₂)

$$= \left(\frac{1}{n} - \frac{1}{N} \right) s_{by}'^2 + \frac{1}{nN} \sum_{i=1}^n u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i} \right) \{ s_{iy}^2 + \hat{R}_i^2 s_{ix}^2 - 2\hat{R}_i s_{ixy} \}$$

Est MSE(T₃)

$$= \left(\frac{1}{n} - \frac{1}{N} \right) \{ s_{by}'^2 + \hat{R}^2 s_{bx}'^2 - 2\hat{R}s'_{bxy} \} + \frac{1}{nN} \sum_{i=1}^n u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i} \right) s_{iy}^2$$

Est MSE(T₄)

$$= \left(\frac{1}{n} - \frac{1}{N} \right) \{ s_{by}'^2 + \hat{R}^2 s_{bx}'^2 - 2\hat{R}s'_{bxy} \} + \frac{1}{nN} \sum_{i=1}^n u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i} \right) d_i^2$$

where $d_i^2 = s_{iy}^2 + \hat{R}_i^2 s_{ix}^2 - 2\hat{R}_i s_{ixy}$

Est MSE (T₅)

$$= \left(\frac{1}{n} - \frac{1}{N} \right) s_{by}'^2 + \frac{1}{nN} \sum_{i=1}^n u_i^2 \left(\frac{1}{m_i} - \frac{1}{M_i} \right) \{ s_{iy}^2 + \hat{R}^2 s_{ix}^2 - 2\hat{R}s_{ixy} \}$$

VII. NUMERICAL ILLUSTRATIONS

For numerical illustration of efficiencies of different estimators, we consider data from 1971 census of India, described below. The population consists of 104 blocks (ssu) divided in to N = 15 wards (fsu) of Berhampur City of Odisha. The number of block (M_i) in 15 wards are 6, 6, 12, 5, 6, 6, 10, 5, 6, 6, 6, 6, 12, 6. The two variables i.e. number of educated females, female population are used as y, x, respectively. For comparison of Mean square error (MSE) of T₀, T₁, T₂, T₃, T₄, T₅, we consider 30% sampling fraction at both the stages. Let the first stage sample size be n= 5 and the sizes of the second stage sample i.e. m_i (i = 1, 2, ..., 15) are assumed to be 2, 2, 4, 2, 2, 2, 3, 2, 3, 3, 2,2, 2, 4 and 3 respectively.

Table 2: Statistical Computations

| i | M _i | m _i | u _i | R _i | (u _i Ȳ _i - Ȳ) ² | (u _i X̄ _i - X̄) ² | (u _i X̄ _i - X̄)(u _i Ȳ _i - Ȳ) | S _{iy} ² | S _{ix} ² | S _{ixy} |
|----|----------------|----------------|----------------|----------------|--|--|--|------------------------------|------------------------------|------------------|
| 1 | 6 | 2 | 0.865 | 0.124 | 4400.504 | 620.313 | 1652.178 | 506.4 | 405.067 | -40 |
| 2 | 6 | 2 | 0.865 | 0.355 | 21.2281 | 503.926 | 103.428 | 1537.066 | 361.1 | 138.8 |
| 3 | 12 | 4 | 1.730 | 0.282 | 2144.551 | 51236.223 | 10482.304 | 1937.174 | 2095 | 676.318 |
| 4 | 5 | 2 | 0.721 | 0.309 | 1618.903 | 9612.665 | 3944.867 | 807.2 | 212.32 | 116.3 |
| 5 | 6 | 2 | 0.865 | 0.435 | 15.246 | 2690.694 | -202.544 | 1538.167 | 169.5 | -176.9 |
| 6 | 6 | 2 | 0.865 | 0.386 | 40.172 | 2330.176 | 305.956 | 2183.867 | 440.267 | 898.134 |
| 7 | 10 | 3 | 1.442 | 0.198 | 681.048 | 6231.397 | -2060.068 | 1114.678 | 328.844 | 76.378 |
| 8 | 5 | 2 | 0.721 | 0.242 | 2390.114 | 6541.897 | 3954.223 | 1934.2 | 1966.0 | 1836.5 |
| 9 | 6 | 3 | 0.865 | 0.355 | 260.599 | 3014.097 | 886.269 | 3097.6 | 2033.2 | -310.2 |
| 10 | 6 | 3 | 0.865 | 0.332 | 680.959 | 4746.080 | 1797.745 | 1108.3 | 269.767 | 257.1 |
| 11 | 6 | 2 | 0.865 | 0.505 | 1428.824 | 297.766 | -652.269 | 2505.866 | 601.9 | 1006 |
| 12 | 6 | 2 | 0.865 | 0.334 | 58.273 | 219.167 | 113.011 | 1007.766 | 488.987 | 589.866 |
| 13 | 6 | 2 | 0.865 | 0.496 | 1342.951 | 219.167 | -542.522 | 3470.4 | 3470.4 | 1017 |
| 14 | 12 | 4 | 1.730 | 0.344 | 5160.784 | 43098.881 | 14913.887 | 1966.697 | 1079.69 | - |
| 15 | 6 | 3 | 0.865 | 0.531 | 2104.662 | 250.088 | -725.501 | 1062.267 | 869.367 | 405.424 |
| | | | | | | | | | | 475.534 |

Calculated Results:

N = 15 n = 5

$\bar{X} = 2888.558$ $\bar{Y} = 99.221$ R = 0.343

$S'_{bx}{}^2 = 9400.895$ $S'_{by}{}^2 = 1596.344$ $S'_{bxy} = 2426.497$

Table 3: Comparison of Mean Square Errors (MSES)

| Estimator | MSE | Efficiency |
|----------------|---------|------------|
| T ₀ | 302.105 | 100 |
| T ₁ | 221.403 | 136.45 |
| T ₂ | 295.624 | 102.19 |
| T ₃ | 227.860 | 132.58 |
| T ₄ | 221.334 | 136.49 |
| T ₅ | 295.693 | 102.17 |

Remarks : For the given illustration, it is observed that

$$MSE(T_4) < MSE(T_1) < MSE(T_3) < MSE(T_2) < MSE(T_5) < MSE(T_0)$$

VIII. CONCLUSION

The ratio estimators considered in a two stage sampling set up have been compared with the usual unbiased estimator without use of auxiliary information, as regards their efficiencies. The numerical illustration shows that T₁ and T₄ have nearly equal efficiency and are more efficient than other competitive estimators.

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