

Impact of Microcredit Programs on Female Headed Households in Jimma Zone, Ethiopia

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Abstract- Microfinance and microcredit is a recent development issue and one of the mechanisms to empower poor people in rural Ethiopia. Microfinance interventions may empower women by increasing their incomes and their control over that income. This study evaluates the impact of microcredit program on female headed households' annual expenditure in rural Ethiopian, Jimma Zone. Data were collected from 1092 female headed households from both participants and non-participants in microcredit program during 2013. Data was analyzed using the PSM technique. PSM results show that participant in microcredit program has a significant and positive impact on households' annual expenditure. Furthermore, the Probit result shows that higher price of egg and sugar makes households more likely to participate in microcredit program. However, more educated households', large land holders and higher income earners are less likely to participate in the program.

Index Terms- Female Headed Households, Microcredit Participation and Propensity Score Matching

I. INTRODUCTION

Many studies often defined the terms microcredit and microfinance interchangeably, however it is important to recognize the distinction between the two. Microcredit defined as the practices of delivering financial loan to poor clients. On the other hand Microfinance is the act of providing these same borrowers with financial services, such as saving and insurance policies. In short, microfinance encompasses the field of microcredit (Sengupta *et.al* 2008).

Microfinance institutions (MFIs) take in a wide range of providers that vary in legal structure, mission, and methodology to offer these financial services for poor clients who do not have access to mainstream banks or other formal financial service providers (Gutu, 2014). Microfinance is a kind of service that serve as the supply of loans, savings, money transfers, insurance, and other financial services to low-income people. Microfinance is a place for the poor and near poor clients to get access to a high quality financial service, which include not just credit but also savings, insurance and fund transfers.

However, Microcredit or known as micro lending is defined as an extremely small loan given to poor people to help them become self employed (Nawai and Shariff, 2010). Microcredit was given to the underprivileged individuals for income-generating activities that will improve the borrowers' living standards. The loan characteristics are, too small, short-term credit (a year or less), no collateral required, weekly repayment, poor borrower and mostly women who are not qualified for a

conventional bank loan. Usually the loan pays high interest rates because of the high cost in running microcredit program and high risk associated with it.

Since collateral is often not required, the effectiveness of the program becomes the main issue for the microcredit institutions to continue providing this services. This is because most of the microcredit institutions are Non- Governmental Organizations (NGOs), that received funds from the government and donors and they are not profit oriented organizations (Ibid).

Beginning in the mid-seventies, savings and credit institutions started extending small loans to groups of poor women in the villages in order to empower them to invest in microenterprises. This form of micro-enterprise credit is based on solidarity group lending where every group member is tasked to ensure the repayment of all members (Gutu, 2014). This would have been one solution to reduce risk associated with this service. However, regarding delivery of financial service, access to women's credit was very limited in Ethiopia. Because of this limited access, the majority of the poor get financial services from informal sources: like money lenders, Iqub¹, Iddr², merchants, friends and relatives etc. The formal financial institutions have not been interested in delivering credit to the poor because of collateral requirements (Abafita. 2003).

Like in other areas of the world, people in Ethiopia are living under poverty. Financial institutions in general, microfinance programs in particular play a crucial role in the empowerment of poor women. However, studies on the impact of microfinance participation by women were extremely limited. Thus, this study has been undertaken to analyze the impact of microfinance participation on female headed households. Furthermore this study tried to understand the institutional and socio-economic factors that affect microfinance program participation in the area and the extent to which this program participation would improve the consumption expenditure of the households.

Problem Statement and Research Question

Benefiting women from microfinance credit schemes is a targeting technique to supplement subsistence level of

¹ Iqub is a regular type of social connection in Ethiopia; basically a group of people together collect equal amount of money and distribute that money term by term for each members. Members of the group used the money to run a business or often to smooth their consumption.

² Iddr is another form of social institution in Ethiopia that basically serves as traditional insurance for those people in the member.

agricultural production and empower them in rural Ethiopia. Micro-finance interventions may empower women by increasing their incomes and their control over that income, enhancing their knowledge and skills in production and trade, and increasing their participation in household decision-making like on expenditure. As a result, social attitudes and perceptions may change, and women's status in the household and community may be enhanced (Kabeer, 1996).

The active participation of women in agriculture is the recent phenomena of rural economic development. Nweze (1995) observed that rural women: are too poor to save, lack ability to organize financial self-help activities and need cheap credit to expand production and income in their farms and non-farm activities. Nwajiuba (1999) put emphasis on the centrality of credit, especially for women farmers to increase their investment in the absence of adequate savings. Credit is a critical input because it can be used to overcome financial obstacles. However, women farmers were perpetually marginalized in the institutionalized credit programmers in Ethiopia.

Many development programs have been extending reasonable amount of credits to rural women since. However, factors contributing to the low economic performance of rural women and the effect of microfinance program participation were not yet studied. In this study, the main objective of propensity score matching was microfinance program participation, the treatment (microfinance participation by female headed households) and potential outcome (annual expenditure by households). The idea was to match those women microfinance participants with that of a control group (non participant women) sharing similar observable characteristics. The mean effect of microfinance participation would calculate as the average difference in expenditure between participants and non participants i.e. the impact would be the change in household annual expenditure as outcome indicator. The use of PSM method was used to answer the question: what would be the total annual expenditure of female headed households had these households not participated in microcredit services?

Empirical Literature

A majority of microfinance programs generally target women often more financially responsible at repaying than men counter parts. Providing and empowering them should be the main policies of microfinance programs (Setboonsarng et. al, 2008). Even though, women are not explicitly excluded from the credit services in developing countries, they have received practically no credit from the formal financial institutions (Ibid). Several reasons were given by researchers in this regard. First, credit services are administered by socio-cultural reasons and women are passive participants in every village. Second, microfinance service is often at the recommendation of the agricultural extension agents whose contacts are primarily with men. Third, formal financial institutions require collateral for their loan since most type of businesses done by poor people especially women are more risky (Ibid). But the property right of most developing countries favored for men as compared to women. Thus women are by default excluded from such microfinance and microcredit programs. Furthermore, lending institutions usually demand a financial guarantee for any loan, of which women are rarely capable.

Most empirical researches on microfinance have been done to assess the repayment performance of borrowers and were considered more or less the same socio-economic characteristics. Berhanu (2005) while studying on the factors that affect loan repayment performances, by employing the Tobit model, variables like land holding size of the family, agro-ecology of the area, total livestock holding, number of years of experience, number of contacts, sources of credit and income from off-farm activities are found to be significant in affecting loan repayment performance. However, the other variables (family size, distance between main road and household residence, purpose of borrowing, loan amount and expenditure for social festivals) are found to be insignificant variables.

Women borrowers in rural areas of Ethiopia used microfinance services for consumption smoothing purposes and they encountered default. They used this income for financing social ceremonies and would not empower them. Assefa (2002) employed a logit model to estimate the effects of hypothesized explanatory variables on the repayment performance of rural women credit beneficiaries in Dire Dawa, Ethiopia. Variables like farm size, annual farm revenue, celebration of social ceremonies, loan diversion, group effect and location of borrowers from lending institution are found to be statistically significant in affecting the loan repayment performance. Retta (2000, cited in Abafita, 2003) employed probit model for loan repayment performance of women fuel wood carriers in Addis Ababa. His finding is that of loan, supervision, suitability of repayment period and other income sources are found to encourage repayment hence reduce the probability of loan default. While educational level is negatively related to loan repayment.

Belay (2002), used a binary Logit model to analyze factors influencing loan repayment performance of rural women. Location of borrowers from lending institution, loan diversion, annual farm revenue and celebration of social ceremonies were highly important in affecting loan repayment performance. In addition Reta (2011) found that age and business types were important in influencing loan repayment performance of the borrower. In addition, gender and business experience of the respondents were found to be significant determinants of loan repayment rate.

Ughomeh *et.al* (2008) investigated the determinants of loan repayment performance among women self-help groups in Bayelsa State, Nigeria. The study revealed that credit was available for agricultural production, processing and petty trading among women farmers. Loan repayment percentage was determined to be 83.73% while percentage default was 17.27%. The estimated regression model indicated that women as household heads, interest rate and household size, negatively and significantly affected the loan repayment performance while price stability of farm proceeds and commitment to self help groups, positively and significantly affect the loan repayment of women farmers in self help groups in the area.

As far as knowledge is concerned few empirical researches have been done on the impact of microfinance by using propensity score matching (PSM) analysis. Setboonsarng *et. al*, (2008) on their study on Microfinance Khushhali Bank, in 2005, used the (PSM) method to address the selectivity bias. They founds that the lending program contributed significantly to

income generation activities such as agricultural production and, in particular, animal raising. However, the impacts on education, health, female empowerment, and so forth were of limited significance.

In addition some other empirical researches confirmed that microfinance participation improve the welfare of participant households. Ghalib et, al (2011), in their study in Pakistan employ a quasi-experimental research design and make use of the data collected by interviewing both borrowers and non-borrower households and control for sample selection biases by using propensity score matching. In this study household access to microfinance reduce poverty. It has been confirmed that microfinance programs had a positive impact on the welfare of participating households.

Arun et,al, (2006), applied propensity score matching method in their to analyze the effect of Micro Finance Institutions (MFIs) on poverty reduction of households in India. The propensity score matching method was employed to estimate the poverty reducing effects of the access to microfinance institutions. Significantly positive effects on the multidimensional poverty indicator suggest the role of microfinance institutions in poverty reduction. It also showed that households in rural areas need loans from microfinance for productive purposes to reduce poverty, while simply accessing to microfinance program is sufficient for urban households to reduce it.

In general, most empirical researches on female headed households have been done on loan the repayment performance. However, few other research outcomes also studied on the

impact of microcredit program participation on rural people though researchers used different variables in their empirical findings. Hence, empirical studies on female headed households are scanty. Therefore, this study has been conducted to contribute in the economic literature on the impact of microcredit program participation on female headed household annual expenditure in Ethiopia.

II. METHODOLOGY

This part presents the research methodology. It describes the data collected and the specific procedures used in evaluating the impact of microcredit program participation on female headed households by using PSM.

Data and Data Collection Procedure

The data used in this study was collected from three microfinance (Eshet, Harbu and Oromia) around Jimma zone, Ethiopia. The data was collected by loan officers of each micro-finance for a special use during 2014. Information was collected from sample women microfinance participants and non-participants during 2013 about their socio-economic characteristics like family resource level (income), distance from market, distance from lending institutions, price of marketable products, etc, and individual characteristics; age, education status and family size was obtain through questionnaires.

Table 1: Description of variables and measurement

Variables	Type	Measurement	Variable label
dist1	continuous	distance/ km	distance from market
dist2	continuous	distance/ km	distance from lending institutions
pcoffee	continuous	local currency	daily household expenditure on coffee ³
psugar	continuous	local currency	price of sugar in the Kebele ⁴
exp	continuous	local currency	total annual expenditure
female	dummy	yes=1, no=0	female participation in the program
age	continuous	year	age of the household head (female)
educ	continuous	year of education	education level of the household head
famz	continuous	number	family size of the household
land	continuous	local measurement ⁵	total land size of the household
income	continuous	local currency	total income of the household
cerem	dummy	yes=1, no=0	social ceremony in the year
peg	continuous	local currency	price of one egg
psalt	continuous	local currency	price of salt in the Kebele

Source: own computation

³ Every day average estimated coffee consumption expenditure of households measured in Ethiopian birr.

⁴ Kebele is the name of small geographical regions (like small village) in Ethiopia.

⁵ In the local areas land is measured by "ttimad"

Empirical Model

In most studies propensity score matching method has been used to evaluate public policies and programs. This method enable researchers to extract information from the sample of microfinance participants (treated) households and a set of matching households that look like the non participant (controlled) households in all relevant pre-intervention characteristics. In other words, PSM matches each adopter household with a non-participant household that has almost the same likelihood of adopting any social programs. The aim of matching is to find the closest comparison group from a sample of nonparticipants to the sample of microfinance service participants.

In this study program participation (Microcredit participation) indicator D_i equals one if individual i participated and zero otherwise. The treatment effect for an individual i written as:

$$\tau_i = Y_i(1) - Y_i(0) \tag{1}$$

The fundamental evaluation problem arises because only one of the potential outcomes is observed for each individual i . The unobserved outcome is called counter- factual outcome.

The average treatment effect (ATT) is that the parameter that received the most attention in evaluation literature, which is defined as:

$$\tau_{ATT} = E(\tau/D = 1) = E[Y(1)/D = 1] - E[Y(0)/D = 1] \tag{2}$$

As the counterfactual mean for those being treated $E[Y(0)/D = 1]$ is not observed, one has to choose a proper substitute for it in order to estimate ATT. Using the mean outcome of untreated individuals $E[Y(0)/D = 0]$ is in non-experimental studies usually not a good idea, because it is most likely that components which determine the treatment decision also determine the outcome variable of interest. Thus, the outcomes of individuals from treatment and comparison group would differ even in the absence of treatment leading to a 'self-selection bias'. For ATT it can be noted as:

$$E[Y(1)/D = 1] - E[Y(0)/D = 0] = \tau_{ATT} + E[Y(0)/D = 1] - E[Y(0)/D = 0] \tag{3}$$

The difference between the left hand side of equation [3] and τ_{ATT} is the so-called 'self-selection bias'. The true parameter τ_{ATT} is identified:

$$E[Y(0)/D = 1] - E[Y(0)/D = 0] = 0 \tag{4}$$

⁶ The treatment effect for an individual i

Thus, in a program evaluation literature the effectiveness of matching estimators as a feasible estimator depends on two fundamental assumptions:

Conditional Independence Assumption [CIA]: This assumption imposes a restriction that choosing to participate in a program is purely random for similar individuals. Given a set of observable covariates (X) which are not affected by treatment, potential outcomes are independent of treatment assignment. This assumption implies that the selection is solely based on observable characteristics, and variables that influence treatment assignment and potential outcomes are simultaneously observed.

$$Y(0), Y(1) \perp D / X, \forall X \tag{5}$$

It should also be clear that conditioning on all relevant covariates is limited in case of a high dimensional vector X. The propensity score $P(D = 1/X) = P(X)$, i.e. the probability for an individual to participate in a treatment given his observed covariates X, is one possible balancing score. The conditional independence assumption (CIA) based on the propensity score (PS) can be written as:

$$Y(0), Y(1) \perp D / P(X), \forall X \tag{6}$$

Common Support: A further assumptions besides conditional independence (CIA) is the common support or overlap condition. The assumption is that P(x) (probabilities) lies between 0 and 1. This restriction implies that the test of the balancing property is performed only on the observations whose propensity score belongs to the common support region of the propensity score of treated and control groups (Becker and Ichino, 2002). Individuals that fall outside the common support region would be excluded in the treatment effect estimation. This is an important condition to guarantee improving the quality of the matching used to estimate the ATT.

$$(Overlap) \quad 0 < P(D = 1/X) < 1 \tag{7}$$

Common support ensures that individuals with the same characteristics have positive probability of being treated or not treated in the program.

Finally the PSM estimator for ATT can be written as:

$$\tau_{ATT}^{PSM} = E_{P(X)/D=1} \{ E[Y(1)/D = 1, P(X)] - E[Y(0)/D = 0, P(X)] \} \tag{8}$$

The explanation of the above PSM estimator is simply the mean difference in outcomes over the common support, appropriately weighted by the propensity score distribution of participants.

Choosing Matching Algorithm

All matching estimators contrast the outcome of a treated individual with outcomes of controlled group members. PSM estimators differ not only in the way the neighborhood for each treated individual is defined and the common support problem is handled, but also with respect to the weights assigned to these neighbors.

Nearest Neighbor Matching: The most straightforward matching estimator is nearest neighbor (NN) matching. The individual from the comparison group is chosen as a matching partner for a treated individual that is closest in terms of propensity score. A problem which is related to NN matching without replacement is that estimates depend on the order in which observations get matched. Hence, when using this approach it should be ensured that ordering is randomly done.

Caliper and Radius Matching: NN matching faces the risk of bad matches, if the closest neighbor is far away. This can be avoided by imposing a tolerance level on the maximum propensity score distance (caliper). To overcome this problem the caliper matching algorithm is another alternative. Caliper matching means that an individual from the comparison group is chosen as a matching partner for a treated individual that lies within a given caliper (propensity score range) and is closest in terms of propensity score (Caliendo and Kopeinig, 2008). Imposing a caliper works in the same direction as allowing for replacement. Bad matches are avoided and hence the matching quality rises. However, if fewer matches can be performed, the variance of the estimates increases.

Stratification and Interval Matching: The idea of stratification matching is to partition the common support of the propensity score into a set of intervals (strata) and to calculate the impact within each interval by taking the mean difference in outcomes between treated and control observations. Clearly, one question to be answered is how many strata should be used in empirical analysis.

Kernel and Local Linear Matching: The matching algorithms discussed so far have in common that only a few observations from the comparison group are used to construct the counterfactual outcome of a treated individual. Kernel matching (KM) and local linear matching (LLM) are non-parametric matching estimators that use weighted averages of all individuals in the control group to construct the counterfactual outcome. Thus, one major advantage of these approaches is the lower variance which is achieved because more information is used. A drawback of these methods is that possibly observations are used that are bad matches. Hence, the proper imposition of the common support condition is of major importance for KM and LLM.

Assessing the Matching Quality

Since we do not condition on all covariates but on the propensity score, it has to be checked if the matching procedure is able to balance the distribution of the relevant variables in both the control and treatment group. Several procedures to do so are discussed in this subsection. These procedures can also, as already mentioned, help in determining which interactions and higher order terms to include in the propensity score specification for a given set of covariates. The basic idea of all approaches is to compare the situation before and after matching

and check if there remain any differences after conditioning on the propensity score. If there are differences, matching on the score is not (completely) successful and remedial measures have to be done, e.g. by including interaction-terms in the estimation of the propensity score. The followings are common criteria to assess the matching qualities.

Standardized Bias: One suitable indicator to assess the distance in marginal distributions of the variables is the standardized bias (SB). For each covariate X it is defined as the difference of sample means in the treated and matched control subsamples as a percentage of the square root of the average of sample variances in both groups. The standardized bias before matching is given by:

$$SB_{before} = 100. \frac{(\bar{X}_1 - \bar{X}_0)}{\sqrt{0.5 \cdot (V_1(X) + V_0(X))}}$$

[9]

The standardized bias after matching is given by:

$$SB_{after} = 100. \frac{(\bar{X}_{1M} - \bar{X}_{0M})}{\sqrt{0.5 \cdot (V_{1M}(X) + V_{0M}(X))}}$$

[10]

Where X_1 (V_1) is the mean (variance) in the treatment group before matching and X_0 (V_0) the analogue for the control group. X_{1M} (V_{1M}) and X_{0M} (V_{0M}) are the corresponding values for the matched samples.

t-test: A similar approach uses a two-sample t-test to check if there are significant differences in covariate means for both groups (Rosenbaum and Rubin, 1985). Before matching differences are expected, but after matching the covariates should be balanced in both groups and hence no significant differences should be found. The t-test might be preferred if the evaluator is concerned with the statistical significance of the results. The shortcoming here is that the bias reduction before and after matching is not clearly visible.

Joint Significance and Pseudo-R²: Additionally, Sianesi (2004) suggests re-estimating the propensity score on the matched sample that is only on participants and matched non-participants and compare the pseudo-R²s before and after matching. The pseudo-R² indicates how well the regressors X explain the participation probability. After matching there should be no systematic differences in the distribution of covariates between both groups and therefore, the pseudo-R² should be fairly low.

Therefore, this study follows and adopts the appropriate empirical methodology. The necessary estimation techniques along with PSM method were used. It describes the average treatment effect; the common support regions; the matching algorithms and the matching quality are presented.

III. RESULTS

Prior to running the Probit regression model to estimate propensity scores, the explanatory variables were checked for the existence of multicollinearity problem. Variance inflation factor

(VIF) was calculated to check the problem of multicollinearity among explanatory variables (*see appendix A*), all explanatory variables are significantly far below 10 (no serious problem of multicollinearity). *STATA 13* version software was used to estimate the PSM model and the effect of microfinance program participation on the annual expenditure of female headed households. The PSM model tried to match 520 controlled female headed households with 572 treated one.

Table2: Simple Comparison of Treated and Controlled Households by Expenditure

Group	Obs	mean	Std. Err.	Std.Dev	[95% Conf.	Interval]
controlled	520	5575.546	180.4809	4115.599	5220.983	5930.109
treated	572	5457.349	167.1918	3998.647	5128.963	5785.735
combined	1092	5513.634	122.6592	4053.325	5272.959	5754.308
diff		118.1971	245.6836		-363.8691	600.2633

Source: Own computation

There is no mean significance difference between the treated and controlled households. From the above table we observed that the comparison and controlled households almost had the balanced number of observations and it makes the analysis more robust.

Estimation of Propensity Scores

The first stage in the propensity score matching is to estimate the probability of being a microcredit or microfinance participant. For this purpose, this study considered variables that influence the likelihood of borrowing from microfinance. The rationale behind this is that, if a variable influences participation but not the outcome, there is no need to control for differences with respect to this variable in the treatment versus the control groups. Likewise, if the variable influences the outcome but not the treatment likelihood, there is no need to control for that variable since the outcome will not significantly differ in the treatment versus the control groups. Variables that affect neither treatment nor the outcome are also clearly unimportant

(Setboonsarng et. al, 2008). Therefore, only those variables that influence both the treatment and the outcome are needed for the matching and are included in the Probit model.

Thus, the Probit regression model was used to estimate the propensity score matching for participant and non-participant female headed households. The dependent variable is binary that indicate households’ participation in microfinance and microcredit services. Results presented in Table 3 below shows the estimated model appears to perform well for the intended matching exercise because the pseudo-R² value is 0.0608. A low R² value shows that program households do not have much distinct characteristics overall (Tolemariam, 2010) and as such finding a good match between program and non-program households becomes easier.

Table 3: Probit Regression Result

female	Coef.	Std. Err.	Z	P>[Z]	[95% Conf.	Interval]
dist1	.0009575	.0157137	0.06	0.951	-.0298408	.0317558
dist2	-.024542	.0375957	-0.65	0.514	-.0982282	.0491442
pcoffee	.1735661	.1355659	1.28	0.200	-.0921383	.4392704
psugar	.0212326	.0101011	2.10	0.036**	.0014348	.0410304
age	.0039881	.003261	1.22	0.221	-.0024034	.0103796
educ	-.0405907	.0124198	-3.27	0.001***	-.0649332	-.0162483
famz	-.0017576	.0190778	-0.09	0.927	-.0391494	.0356341
land	-.001512	.0004622	-3.27	0.001***	-.0024178	-.0006061
income	-4.66e-07	2.35e-07	-1.98	0.048**	-9.28e-07	-4.43e-09
cerem	-.2026552	.1088702	-1.86	0.063*	-.416037	.0107266
peg	.3309907	.1347578	2.46	0.014**	.0668702	.5951112
psalt	.034816	.0490617	0.71	0.478	-.0613432	.1309752
-cons	-1.905136	.8248024	-2.31	0.021**	-3.521718	-.2885526

<i>Number of obs</i>	1092
<i>LR chi2 (12)</i>	91.82
<i>Prob > chi2</i>	0.0000
<i>Pseudo R2</i>	0.0608

***, ** and * is significant at % 1, 5% and 10% level of significance respectively

Source: Own computation

Looking into the estimated coefficients above, the results indicate that program participation is influenced by six explanatory variables. Price of sugar, education level of the household, land size, annual income, social ceremony per year and price of egg variables are which affect the participation of the household to the microfinance program. In this probit regression result some variables like price of egg and price of sugar are makes households more likely to participate in microfinance program. On the contrary, more educated households', those households with large land size and higher

income households are less likely to participate in the program. In addition, social ceremony celebrators are less likely to participate in the program (at 10% level of significance).

Common Support

The final number of blocks in this model is 4. This number of blocks insures that the mean propensity score is not different from treated and controls in each blocks. The region of common support is [0.05302391, 0.77618416] implying that the two groups share the same characteristics in these brackets.

Table 4: Region of Common Support

Estimated propensity score				
	percentiles	smallest		
1%	.1211147	.0530239		
5%	.260821	.0672431		
10%	.3489146	.0705541	obs	1082
25%	.4727205	.0724818	Sum of wgt.	1082
50%	.5525003		mean	.529477
		<u>Largest</u>	Std. Dev	.1287829
75%	.6179895	.7447806		
90%	.667089	.7588135	Variance	.016585
95%	.689535	.7757975	Skewness	-1.162977
99%	.7258213	.7761842	Kurtosis	4.4563

Source: Own computation

Matching Algorithm

Alternative matching estimators were used in matching the treatment and control households in the common support region. The following result (Table 5) shows that female microfinance participation does have a significant impact on household per capita expenditure by the nearest-neighborhood matching method at 5 percent level of significance ($t = 2.242$). The average treatment effect of the treated (ATT) on household expenditure for female headed program participation was 611.699. Participants are on average more annually expend Birr 611.699 as compared to non participants.

The average treatment effect using stratification matching result that follows shows 555.598 more annual expenditure of female head household because of program participation. The impact is significant at 5 percent level ($t = 2.646$). Furthermore, the ATT using radius matching result shows an increased impact (721.797) more household expenditure significantly ($t = 3.097$) of women's microcredit participation on per capita expenditure. The number of individuals matched more or less consistent in all matching algorithms. Looking in to the average treatment effect using kernel matching result is consistent with earlier findings. The robustness of the result of this matching algorithm, women's participation increases per capita expenditure by 652.645 at a 5 percent significance level. This result was quite similar with nearest neighbor, radius and stratification matching methods.

Table5: Results from Different Matching Algorithm

Matching type ⁷	n. treat. ⁸	n. contr. ⁹	ATT ¹⁰	Std. Err.	t
NN	572	275	611.699	272.874	2.242**
Stratification	572	510	555.598	210.014	2.646**
Radius	501	399	721.797	233.072	3.097**
Kernel ¹¹	572	510	652.645	229.275	2.847**

** Significant at 5% level

Source: Own computation

Thus, all matching algorithms used above gives similar results. The maximum amount of treated households considered in estimation of ATT was 572 whereas the minimum was 501. The minimum amount of controlled among these matching algorithms was 275 and the maximum 510. The maximum average expenditure because of credit program participation in these matching algorithms was 721.797 while the minimum was 555.598. All results were statistically significant at five percent.

Matching Quality

Table 6: Matching Quality

variables	Unmatched	mean		%reduct		t-tast		V(T)	
		Matched	treated	control	biase	[Biase]	t	p>[t]	V(C)
dist1	U		12.822	12.904	-2.4		-0.39	0.693	0.87
	M		12.827	12.743	2.5	-3.5	0.43	0.668	0.97
dist2	U		8.9783	8.9619	1.5		0.25	0.799	1.00
	M		8.9802	8.9727	0.7	54.6	0.12	0.906	0.99
pcoffee	U		3.9721	3.9429	7.8		1.29	0.199	1.15
	M		3.9694	3.9643	1.4	82.4	0.23	0.816	1.16
Psugar	U		41.666	41.123	13.4		2.21	0.027	1.06
	M		41.652	41.677	-0.6	95.5	-0.11	0.916	1.22*
age	U		46.164	45.95	1.7		0.28	0.780	0.75*
	M		46.083	45.761	2.5	-50.2	0.43	0.670	0.74*
educ	U		1.729	2.9385	-35.2		-5.84	0.000	0.62*
	M		1.7381	1.7276	0.3	99.1	0.06	0.952	1.05
famz	U		5.1923	5.3827	-8.7		-1.45	0.148	0.79*
	M		5.188	5.1769	0.5	94.2	0.09	0.931	0.82*
land	U		42.965	115.79	-35.1		-5.91	0.000	0.08*
	M		43.181	46.988	-1.8	94.8	-0.70	0.483	0.64*
income	U		75744	2.5e+05	-19.3		-3.27	0.001	0.02*
	M		76049	84899	-1.0	94.8	0.70	0.486	0.51*
cerem	U		.81469	.84615	-8.4		-1.38	0.167	.
	M		.81898	.81291	1.6	80.7	0.26	0.792	.
peg	U		1.5864	1.5467	12.0		1.98	0.048	1.03
	M		1.5858	1.5805	1.6	86.4	0.28	0.781	1.11
psalt	U		9.5073	9.4668	4.8		0.80	0.426	0.86
	M		9.5052	9.5358	-3.6	24.4	-0.61	0.541	0.86

⁷ All matching algorithms; nearest neighbor (NN), Stratification, Radius and Kernel matching are used.

⁸ means number of treated observations

⁹ number of controlled households

¹⁰ Average treatment on the treated

¹¹ A type of algorism used in this study that tested the robustness of the results

Source: Own computation

After choosing the best performing matching algorithm the next task is to check the balancing of propensity and covariates. The main purpose of the propensity score estimation is not to obtain a precise prediction of selection into treatment, but rather to balance the distributions of relevant variables in both groups. The balancing powers of the estimations are ascertained by considering different test methods such as the reduction in the mean standardized bias between the matched and unmatched households, equality of means using *t*-test and chi-square test for joint significance for the variables used.

The mean standardized bias before and after matching are shown in Table 6 with the total bias reduction obtained by the matching procedure. In all cases, it is evident that sample differences in the unmatched data significantly exceed those in the samples of matched cases. The process of matching thus creates a high degree of covariate balance between the treatment and control samples that are ready to use in the estimation procedure. Similarly, *t*-values in the same table show that before matching five of chosen variables exhibited statistically significant differences while after matching all of the covariates are balanced.

The low pseudo-R² and the insignificant likelihood ratio tests support the hypothesis that both groups have the same distribution in covariates X after matching (see Table 7). This result clearly shows that the matching procedure is able to balance the characteristics in the treated and the matched comparison groups. We, therefore, used these results to evaluate the effect of microcredit participation of households having similar observed characteristics. This allowed us to compare observed outcomes for participants with those of a comparison groups sharing a common support.

Table 7: Matching Quality with Pseudo R²

Sample	Ps R2	LR chi2	p>chi2
Unmatched	0.061	91.82	0.000
Matched	0.001	2.12	0.999

Source: Own computation

IV. CONCLUSION

In this study data from 1092 female headed households from Jimma zone during 2013 were analyzed by *STATA 13* version software. Hence, the study has applied a propensity score matching technique which has become the most widely applied non-experimental tool for impact evaluation of social programs and policies. The main research question of the study was what would have happened to an outcome of interest had the microcredit program not been in place. Answering this question requires observing outcomes with-and-without the program for the same household. However, it is impossible to observe the same object in two states simultaneously.

The result of this study confirms that microcredit participation was significantly affected by six variables; Price of sugar, education level of the household, land size of the household, annual income, social ceremony per year and price of

egg are variables that affect the program participation. Price of sugar and price of egg makes households more likely to participate in the microfinance program. However, more educated households, those households with large land size and higher income households are less likely to participate in the microfinance program. This is the reason why in rural part of Ethiopia illiterates are more poor than literates and hence they required more income from credit institutions. Large land size is another source of income in the study area. Therefore, households with large land holdings and higher income earners are less demander of microcredit services.

The propensity score matching based on different matching algorithms have resulted in different number of participant households to be matched with non-participant households after discarding households whose values were out of common support region. The concluding results based on PSM then indicate that there are significant differences in annual expenditure of households between treatment and comparison households, which could be attributed to the participation of microfinance program. Thus microfinance program has positive and significant effect on female headed households' annual expenditure.

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Appendix

Appendix A: Test of Multicollinearity among Explanatory Variable

Variable	VIF	1/VIF
land	3.32	0.300988
income	3.16	0.316894
dist1	1.86	0.538146
Pcoffee	1.68	0.593643
peg	1.30	0.767529
famz	1.12	0.88999
educ	1.12	0.894867
psalt	1.10	0.906191
age	1.10	0.912496
Psugar	1.10	0.912882
cerem	1.09	0.91394
dist2	1.09	0.914741
female	1.07	0.932641

Appendix B: Common support graph

